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GENERALIZED JENSEN DIFFERENCE BASED ON ENTROPY FUNCTIONS¹

PRASANNA K. SAHOO, A. K. C. WONG

This paper concerns an upper bound for a diversity measure between subpopulations (also called the generalized Jensen difference) based on entropy functions. We show that the diversity measure between the subpopulations with given a priori probabilities and induced by either the Shannon entropy or the entropy of degree α can never exceed the corresponding entropies of the a priori probabilities. Through this bound we prove a conjecture of Wong and You [25] in affirmative and suggest a new definition for the index of diversity based on entropy functions. An upper bound for the second order generalized Jensen difference based on entropy functions is also obtained in this paper.

1. INTRODUCTION

Consider a set of populations $\{\pi_k\}$, where the individuals of each population π_k are characterized by a set of measurements A in a measurable space (Ω, \mathcal{B}) . The set Ω is the sample space and \mathcal{B} is a σ -algebra of subsets of Ω . The discrete probability distribution function of A in π_k is denoted by P_k which belongs to an n -dimensional simplex $A_n := \{P = (p_1, p_2, \dots, p_n) \mid 0 \leq p_k \leq 1, \sum_{k=1}^n p_k = 1\}$. Let A_n^0 be the interior of the n -dimensional simplex A_n , that is, $A_n^0 = \{P = (p_1, p_2, \dots, p_n) \mid 0 < p_k < 1, \sum_{k=1}^n p_k = 1\}$. Notice that A_n and A_n^0 are convex subsets of \mathbb{R}^n . A real valued function $D: A_n \rightarrow \mathbb{R}$ is called a diversity measure on A_n if it satisfies the following conditions:

- (a) $D(P) \geq 0$ for all $P \in A_n$,
- (b) $D(P) = 0$ if, and only if, $p_i = 1$ for some i and $p_j = 0$ for all $j \neq i$,
- (c) D is a concave function on A_n , that is, if $\lambda_1, \lambda_2, \dots, \lambda_m$ are nonnegative real numbers such that $\sum_{i=1}^m \lambda_i = 1$ and $P_1, P_2, \dots, P_m \in A_n$, then

$$(1.1) \quad D\left(\sum_{k=1}^m \lambda_k P_k\right) \geq \sum_{k=1}^m \lambda_k D(P_k).$$

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While the conditions (a) and (b) are natural, the condition (c) was motivated by the consideration that the diversity in a mixture of populations should not be smaller than the average of the diversities within the component populations. Some examples of the diversity function defined on the set Δ_n are the followings (Nayak [14] and [15]):

$$(1.2) \quad H(P) = - \sum_{k=1}^n p_k \log_2 p_k$$

$$(1.3) \quad H_\alpha(P) = \frac{1}{2^{1-\alpha} - 1} \left(\sum_{k=1}^n p_k^\alpha - 1 \right), \quad \alpha > 0, \quad \alpha \neq 1,$$

$$(1.4) \quad H^\alpha(P) = \frac{1}{1-\alpha} \log_2 \left(\sum_{k=1}^n p_k^\alpha \right), \quad \alpha \in]0, 1[,$$

$$(1.5) \quad S(P) = 1 - \sum_{k=1}^n p_k^2,$$

$$(1.6) \quad B(P) = \frac{1}{1-2^{\gamma-1}} \left[1 - \left(\sum_{k=1}^n p_k^{1/\gamma} \right)^\gamma \right], \quad \gamma > 0, \quad \gamma \neq 1,$$

$$(1.7) \quad R(P) = PMP^T,$$

where M is an $n \times n$ matrix such that PMP^T satisfies conditions (a), (b) and (c). The first three diversity measures are the Shannon entropy [22], the entropy of degree α [10], and the entropy of order α [19], respectively. The fourth diversity measure is a special case of the entropy of degree α and is known as the Gini-Simpson index in statistics. The fifth diversity measure in the example is the γ -entropy introduced by Arimoto [1] and axiomatically characterized by Behara and Chawla [2]. The diversity measures (1.3), (1.4) and (1.6) are entropies which depend on a parameter. In applications, one expects that these should be monotonically increasing functions of the parameter. As for the entropy of degree α and the entropy of order α , it is well known that they are monotonically increasing functions of α . Regarding the monotonicity of (1.6), it was conjectured that the γ -entropy is also an increasing function of the parameter γ (see Nayak [14]). Recently, Sahoo [20] has shown that this conjecture is true. The last diversity measure was suggested by Rao [18] and it is often referred to as the Rao's quadratic entropy. The Shannon entropy, the entropy of degree α , the entropy of order α , the Gini-Simpson index, and the γ -entropy are all invariant under permutation of their arguments. As a result, they all have certain flaws. In spite of these flaws, entropic diversity measures are widely used in genetics, sociology, statistics [14, 15, 24], anthropology, and recently in pattern recognition [3, 12, 25].

Given a set of probabilities measures $P_1, P_2, \dots, P_m \in \Delta_n$ and $X = (x_1, x_2, \dots, x_m) \in \Delta_m$, and a concave function D defined on Δ_n , the following decomposition was suggested by Rao ([17] and [18]).

$$(1.8) \quad D\left(\sum_{k=1}^m x_k P_k\right) = \sum_{k=1}^m x_k D(P_k) + J_D(\{P_i\}, \{x_i\}).$$

The second term, $J_D(\{P_i\}, \{x_i\})$, in (1.8) is called the generalized Jensen difference in Burbea and Rao [4] of the function D . If P_1, P_2, \dots, P_m are probability distributions in m subpopulations with a priori probabilities (or a priori weights) $x_1, x_2, \dots, \dots, x_m$, then the term $D(\sum_{k=1}^m x_k P_k)$ is known as the total diversity whereas the quantity $\sum_{k=1}^m x_k D(P_k)$ is the average diversity within subpopulations. Thus, $J_D(\{P_i\}, \{x_i\})$ may be interpreted as the diversity between the subpopulations. In biological works, J_D when D is the Shannon entropy, is defined to be the information radius on probability distributions P_1, P_2, \dots, P_m (see Sibson [23]) and some applications of this concept to cluster analysis are discussed in Sibson [23] and Jardine and Sibson [11]. The ratio

$$G_D = \frac{J_D(\{P_i\}, \{x_i\})}{J_D(\{P_i\}, \{x_i\}) + \sum_{k=1}^m x_k D(P_k)}$$

is the index of diversity induced by the functional D between subpopulations compared to the total. The index G_D has been used widely in genetics by Lewontin [13], Nei [16] and Chakraborty [6] when D is the Shannon entropy. For convenience, we will refer $J_D(\{P_i\}, \{x_i\})$ to as the diversity (between the subpopulations) induced by the function D .

This paper is organized as follows: In Section 2, we prove an upper bound for J_D when D is the Shannon entropy. Section 3 presents an upper bound for J_D assuming D to be entropy of degree α . In Section 4, some inequalities regarding the generalized Jensen difference are proven. In Section 5, we propose a new index of diversity.

2. UPPER BOUND FOR J_D INDUCED BY THE SHANNON ENTROPY

In this section, we prove the following theorem.

Theorem 1. Let $P_i = (p_{i1}, p_{i2}, \dots, p_{in}) \in \Delta_n^0$ for $i = 1, 2, \dots, m$ be the m complete probability distributions. Let $X = (x_1, x_2, \dots, x_m) \in \Delta_m^0$, then inequality

$$(2.1) \quad J_H(\{P_i\}, \{x_i\}) < H(X)$$

holds, where H is the Shannon entropy.

Proof. Using (1.8) and the form of the Shannon entropy in (1.2), one obtains

$$(2.2) \quad J_H(\{P_i\}, \{x_i\}) - H(X) = \sum_{j=1}^m \sum_{k=1}^n p_{jk} \log_2 \frac{p_{jk}}{r_k} + \sum_{i=1}^m x_i \log_2 x_i,$$

where $r_k = \sum_{i=1}^m x_i p_{ik}$ ($k = 1, 2, \dots, n$). Since $\sum_{k=1}^n p_{ik} = 1$ for all $i = 1, 2, \dots, m$, there-

fore (2.2) can be rewritten as

$$(2.3) \quad J_H(\{P_i\}, \{x_i\}) - H(X) = \sum_{i=1}^m \sum_{k=1}^n x_i p_{ik} \log_2 \frac{x_i p_{ik}}{r_k}.$$

The use of the inequality

$$\log_2 x \leq (x - 1) \log_2 e$$

(with equality if, and only if, $x = 1$) in (2.3) results in the following:

$$(2.4) \quad J_H(\{P_i\}, \{x_i\}) - H(X) < \sum_{i=1}^m \sum_{k=1}^n x_i p_{ik} \left(\frac{x_i p_{ik}}{r_k} - 1 \right) \log_2 e.$$

This strict inequality is due to the fact that $P_i \in \mathcal{A}_n^0$ and $X \in \mathcal{A}_m^0$. Some algebraic simplifications of (2.4) yield

$$J_H(\{P_i\}, \{x_i\}) - H(X) < \log_2 e \sum_{i=1}^m \sum_{k=1}^n \frac{(x_i^2 p_{ik}^2 - x_i p_{ik} r_k)}{r_k}.$$

Some further simplifications of the above inequality lead to

$$(2.5) \quad J_H(\{P_i\}, \{x_i\}) - H(X) < \log_2 (e^{-1}) \sum_{k=1}^n \frac{1}{r_k} \sum_{s=1}^n \sum_{t \neq s} x_s x_t p_{sk} p_{tk},$$

where $\sum_{t \neq s}$ denotes the summation over all t except $t = s$. Since x_i and p_{ik} are strictly greater than zero, we have

$$(2.6) \quad \sum_{k=1}^n \frac{1}{r_k} \sum_{s=1}^n \sum_{t \neq s} x_s x_t p_{sk} p_{tk} > 0.$$

Hence, from (2.6) and (2.5) one obtains

$$J_H(\{P_i\}, \{x_i\}) < H(X).$$

This completes the proof of the theorem. \square

Remark 1. The proof of Theorem 1 can also be obtained using the analytic and algebraic properties of mutual information, since $J_H(\{P_i\}, \{x_i\})$ (see Gallager [9] p. 90) is the mutual information associated with the m -input symbol distribution and the so called channel matrix $(p_{ij})_{m \times n}$. However the proof given in this paper is straightforward and does not use the properties of mutual information.

The proof of the following corollary is similar to that of Theorem 1 and is therefore omitted.

Corollary 1. Let $P_i \in \mathcal{A}_n$ ($i = 1, 2, \dots, m$) and $X \in \mathcal{A}_m$. Then the following inequality holds

$$J_H(\{P_i\}, \{x_i\}) \leq H(X)$$

with equality if, and only if, $x_j = 1$ for some j and $x_i = 0$ for all $i \neq j$.

3. UPPER BOUND FOR J_D INDUCED BY THE ENTROPY OF DEGREE α

In this section, we shall derive an upper bound for the diversity (between the sub-population) induced by the entropy of degree α . We denote the entropy of degree α by H_α . Notice that the definition of H_α is valid only for all $\alpha (>0)$ and $\alpha \neq 1$. When $\alpha = 1$, we define H_α through the limit $\alpha \rightarrow 1$. That is,

$$(3.1) \quad H_1 := \lim_{\alpha \rightarrow 1} H_\alpha.$$

It is not difficult to prove, using l'Hopital rule, that the right hand side of (3.1) is the Shannon entropy H .

Theorem 2. Let $P_1, P_2, \dots, P_m \in \mathcal{A}_n^0$ and $X \in \mathcal{A}_m^0$, then $J_{H_\alpha}(\{P_i\}, \{x_i\})$ satisfies the inequality

$$(3.2) \quad J_{H_\alpha}(\{P_i\}, \{x_i\}) < H_\alpha(X), \quad \alpha \in \mathbb{R}_+,$$

where H_α is the entropy of degree α .

Proof. Let $\alpha > 0$ and consider the following

$$\sum_{j=1}^n [(\sum_{i=1}^m x_i p_{ij})^\alpha - \sum_{i=1}^m (x_i p_{ij})^\alpha] - \sum_{i=1}^m x_i^\alpha + 1.$$

Since $\sum_{i=1}^m x_i = 1 = \sum_{j=1}^n p_{ij}$, the above expression can be written as

$$\sum_{j=1}^n [(\sum_{i=1}^m x_i p_{ij})^\alpha - \sum_{i=1}^m (x_i p_{ij}^\alpha + x_i^\alpha p_{ij} - x_i p_{ij})].$$

Let

$$(3.3) \quad \Phi_j(\alpha) := (\sum_{i=1}^m x_i p_{ij})^\alpha - \sum_{i=1}^m (x_i p_{ij}^\alpha + x_i^\alpha p_{ij} - x_i p_{ij}).$$

At this point we would like to prove the following inequalities

$$(3.4) \quad \sum_{j=1}^n \Phi_j(\alpha) \begin{cases} < 0 & \text{for } \alpha \in]0, 1[\\ > 0 & \text{for } \alpha \in]1, \infty[. \end{cases}$$

First consider the case, when $\alpha \in]1, \infty[$. Since $x_i, p_{ij} \in]0, 1[$, therefore $(\sum_{i=1}^m x_i p_{ij})^\alpha > \sum_{i=1}^m x_i^\alpha p_{ij}^\alpha$. Thus, in view of this, (3.3) becomes

$$(3.5) \quad \Phi_j(\alpha) > \sum_{i=1}^m [x_i^\alpha p_{ij}^\alpha - x_i p_{ij}^\alpha + x_i p_{ij} - x_i^\alpha p_{ij}].$$

Rewriting (3.5), we have

$$(3.6) \quad \Phi_j(\alpha) > \sum_{i=1}^m (p_{ij} - p_{ij}^\alpha)(x_i - x_i^\alpha).$$

Since p_{ij} and x_i are in $]0, 1[$, therefore $(p_{ij} - p_{ij}^\alpha) > 0$ if $\alpha > 1$. Similarly we get $(x_i - x_i^\alpha) > 0$. Hence from (3.6), we obtained the second half of the inequality (3.4).

Now consider the case when $0 < \alpha < 1$. Here, the inequality (3.6) will reverse its sign, that is

$$(3.7) \quad \Phi_j(\alpha) < \sum_{i=1}^m (p_{ij} - p_{ij}^\alpha)(x_i - x_i^\alpha),$$

since $(\sum_{i=1}^m x_i p_{ij})^\alpha < \sum_{i=1}^m x_i^\alpha p_{ij}^\alpha$ for $\alpha \in]0, 1[$. In this case $(p_{ij} - p_{ij}^\alpha)$ and $(x_i - x_i^\alpha)$ are strictly less than zero. Thus, by (3.7) the first half of the inequality (3.4) is proven.

By using the definition of H_α and the definition of $J_{H_\alpha}(\{P_i\}, \{x_i\})$, we obtain

$$(3.8) \quad J_{H_\alpha}(\{P_i\}, \{x_i\}) - H_\alpha(X) = (2^{1-\alpha} - 1)^{-1} \sum_{j=1}^n \Phi_j(\alpha).$$

Note that if $\alpha > 1$ then $(2^{1-\alpha} - 1)^{-1} < 0$. Thus, using (3.4) in (3.8) we obtain (3.2). Again, if $0 < \alpha < 1$, then $(2^{1-\alpha} - 1)^{-1} > 0$ and (3.2) follows. The case, when $\alpha = 1$ can be obtained from Theorem 1 since $\lim_{\alpha \rightarrow 1} J_{H_\alpha}(\{P_i\}, \{x_i\}) = J_H(\{P_i\}, \{x_i\})$. This completes the proof of the theorem. \square

Remark 2. Daróczy [7] defined channel capacity of degree α using the quantity $I_\alpha(X, Y) := H_\alpha(Y) - H_\alpha(Y|X)$, where Y is the distribution associated with the mixture, that is, $\sum_{k=1}^m x_k P_k$ and $H_\alpha(Y|X) = \sum_{k=1}^m x_k^\alpha H_\alpha(P_k)$. Thus, $J_{H_\alpha}(\{P_i\}, \{x_i\})$ is not the mutual information of Daróczy and hence one cannot use the properties of mutual information of degree α to obtain another proof of Corollary 1 unlike the case in Theorem 1. For definition and some applications of $H_\alpha(Y|X)$, the conditional entropy of degree α , refer to El-Sayed [8] and Sahoo [21].

The following corollary can be proven by mimicking the proof of Theorem 2.

Corollary 2. Let $P_i \in \mathcal{A}_n$ ($i = 1, 2, \dots, m$) and $X \in \mathcal{A}_m$, then the inequality

$$J_{H_\alpha}(\{P_i\}, \{x_i\}) \leq H_\alpha(X)$$

holds, with equality if, and only if, $x_j = 1$ for some j and $x_i = 0$ for all $i \neq j$.

4. UPPER BOUND FOR SECOND ORDER GENERALIZED JENSEN DIFFERENCE

The concept of the generalized Jensen difference was extended to higher order in [5] by Burbea and Rao. In view of this extension, the Jensen difference J_D defined through (1.8), is the Jensen difference of order one. In this section, we derive an upper bound for the second order Jensen difference based on the Shannon entropy or the entropy of degree α .

Consider the set of n -ary complete probability distributions $\{P_{i_1 i_2} \in \mathcal{A}_n \mid i_1 = 1, 2, \dots, m_1; i_2 = 1, 2, \dots, m_2\}$ index by combinations of the levels i_1 and i_2 of two factors with independent a priori distributions $X^{(j)}$ in \mathcal{A}_{m_j} , where $X^{(j)} = (x_1^{(j)}, x_2^{(j)}, \dots, x_{m_j}^{(j)})$, $j = 1, 2$. The conditional distributions subject to individual

levels of i_1 and i_2 and the unconditional distributions are

$$(4.1) \quad P_{i_1} = \sum_{i_2=1}^{m_2} x_{i_2}^{(2)} P_{i_1 i_2},$$

$$(4.2) \quad P_{i_2} = \sum_{i_1=1}^{m_1} x_{i_1}^{(1)} P_{i_1 i_2},$$

and

$$(4.3) \quad P = \sum_{i_1=1}^{m_1} \sum_{i_2=1}^{m_2} x_{i_1}^{(1)} x_{i_2}^{(2)} P_{i_1 i_2}.$$

The conditional Jensen difference between the levels of $X^{(1)}$ for a given level i_2 of $X^{(2)}$ is, given by (see Burbea and Rao [5])

$$(4.4) \quad J_D^{(1, i_2)} = D\left(\sum_{i_1=1}^{m_1} x_{i_1}^{(1)} P_{i_1 i_2}\right) - \sum_{i_1=1}^{m_1} x_{i_1}^{(1)} D(P_{i_1 i_2}).$$

The Jensen difference between the levels of $X^{(1)}$ averaged over all the levels of $X^{(2)}$ is

$$(4.5) \quad J_D^{(1, \bar{2})} = D(P) - \sum_{i_1=1}^{m_1} D(P_{i_1}).$$

The second order generalized Jensen difference (or interaction between factor 1 and 2) is defined in [5] as

$$(4.6) \quad J_D^{(1, 2)} = \sum_{i_2=1}^{m_2} x_{i_2}^{(2)} J_D^{(1, i_2)} - J_D^{(1, \bar{2})}.$$

In [5], Burbea and Rao studied the convexity of higher order generalized Jensen difference. Regarding second order generalized Jensen difference the followings

- (a) $J_D^{(1, 2)} = J_D^{(2, 1)}$,
- (b) $J_D^{(1, 2)} \geq 0$ if and only $J_D \geq 0$,

are true. For details readers should refer to [5]. Now we prove the following lemma.

Lemma 1. If D is a concave function, then the inequality

$$(4.7) \quad J_D^{(1, 2)} \leq \sum_{i_1=1}^{m_1} x_{i_1}^{(1)} \left[D\left(\sum_{i_2=1}^{m_2} x_{i_2}^{(2)} P_{i_1 i_2}\right) - \sum_{i_2=1}^{m_2} x_{i_2}^{(2)} D(P_{i_1 i_2}) \right] + \\ + \sum_{i_2=1}^{m_2} x_{i_2}^{(2)} \left[D\left(\sum_{i_1=1}^{m_1} x_{i_1}^{(1)} P_{i_1 i_2}\right) - \sum_{i_1=1}^{m_1} x_{i_1}^{(1)} D(P_{i_1 i_2}) \right]$$

holds, where $P_{i_1 i_2} \in \mathcal{A}_n$ ($i_1 = 1, 2, \dots, m_1$; $i_2 = 1, 2, \dots, m_2$) and $X^{(j)} \in \mathcal{A}_m$ ($j = 1, 2$).

Proof. By (4.4) and (4.3), (4.5) becomes

$$(4.8) \quad J_D^{(1, 2)} = \sum_{i_2=1}^{m_2} x_{i_2}^{(2)} \left[D\left(\sum_{i_1=1}^{m_1} x_{i_1}^{(1)} P_{i_1 i_2}\right) - \sum_{i_1=1}^{m_1} x_{i_1}^{(1)} D(P_{i_1 i_2}) \right] + \\ + \sum_{i_1=1}^{m_1} x_{i_1}^{(1)} \left[D\left(\sum_{i_2=1}^{m_2} x_{i_2}^{(2)} P_{i_1 i_2}\right) - D\left(\sum_{i_1=1}^{m_1} \sum_{i_2=1}^{m_2} x_{i_1}^{(1)} x_{i_2}^{(2)} P_{i_1 i_2}\right) \right].$$

Since D is a concave function,

$$(4.9) \quad D\left(\sum_{i_1=1}^{m_1} \sum_{i_2=1}^{m_2} x_{i_1}^{(1)} x_{i_2}^{(2)} P_{i_1 i_2}\right) \geq \sum_{i_1=1}^{m_1} \sum_{i_2=1}^{m_2} x_{i_1}^{(1)} x_{i_2}^{(2)} D(P_{i_1 i_2}).$$

In view of (4.9), (4.8) yields (4.7). This completes the proof of the lemma. \square

Now we proceed to determine an upper bound for $J_D^{(1,2)}$ when D is either the Shannon entropy or the entropy of degree α . We denote $J_{H_\alpha}^{(1,2)}$ as the second order generalized Jensen difference based on the entropy of degree α .

Theorem 3. Let $P_{i_1 i_2} \in \mathcal{A}_n$ ($i_1 = 1, 2, \dots, m_1$; $i_2 = 1, 2, \dots, m_2$) and $X^{(j)} \in \mathcal{A}_{m_j}$ ($j = 1, 2$). Then the inequality

$$(4.10) \quad J_{H_\alpha}^{(1,2)} \leq \sum_{j=1}^2 H_\alpha(X^{(j)})$$

holds for all α in \mathbb{R}_+ .

Proof. Since for α in $\mathbb{R}_+ - \{1\}$, H_α is a concave function of the probabilities, using Lemma 1 we obtain

$$(4.11) \quad \begin{aligned} J_{H_\alpha}^{(1,2)} &\leq \sum_{i_2=1}^{m_2} x_{i_2}^{(2)} \left[H_\alpha\left(\sum_{i_1=1}^{m_1} x_{i_1}^{(1)} P_{i_1 i_2}\right) - \sum_{i_1=1}^{m_1} x_{i_1}^{(1)} H_\alpha(P_{i_1 i_2}) \right] + \\ &\quad + \sum_{i_1=1}^{m_1} x_{i_1}^{(1)} \left[H_\alpha\left(\sum_{i_2=1}^{m_2} x_{i_2}^{(2)} P_{i_1 i_2}\right) - \sum_{i_2=1}^{m_2} x_{i_2}^{(2)} H_\alpha(P_{i_1 i_2}) \right]. \end{aligned}$$

Now Corollary 2 in (4.11) yields

$$(4.12) \quad J_{H_\alpha}^{(1,2)} \leq \sum_{i_2=1}^{m_2} x_{i_2}^{(2)} H_\alpha(X^{(1)}) + \sum_{i_1=1}^{m_1} x_{i_1}^{(1)} H_\alpha(X^{(2)}).$$

Using the fact $\sum_{i_1=1}^{m_1} x_{i_1}^{(1)} = 1$ and $\sum_{i_2=1}^{m_2} x_{i_2}^{(2)} = 1$, we get (4.10) from (4.12). If $\alpha = 1$, by Corollary 1 we again obtain (4.10). This completes the proof of the theorem. \square

5. COMMENTS

By maximality of the Shannon entropy, we obtain the following inequality, an obvious consequence of Corollary 1:

$$(5.1) \quad J_H(\{P_i\}, \{x_i\}) \leq \log_2 m,$$

where $P_1, P_2, \dots, P_m \in \mathcal{A}_n$ and $X \in \mathcal{A}_m$.

In the context of structural pattern recognition, Wong and You [25] observed the necessity of an upper bound for J_D when D is the Shannon entropy. They conjectured on the basis of computer simulation that, for all $x \in [0, 1[$ and for every $P_1, P_2 \in \mathcal{A}_n$, the quantity $H(xP_1 + (1-x)P_2) - xH(P_1) - (1-x)H(P_2)$, which they call 'increment of entropy', is bounded from above by 1 (see [25, p. 604]). From (5.1), it is quite obvious that the conjecture is true.

Since D is a concave function, by Jensen inequality for concave function, we see from (1.8), that $J_D(\{P_i\}, \{x_i\})$ is nonnegative. Thus, the 'increment of entropy' (which is equivalent to the mutual information associated with a 2-symbol source and channel matrix (p_{ij})) can be used as an index of diversity. This index is successfully used in the comparison of random graphs in pattern recognition (see Wong and You [25]). Similar to the index of diversity G_D one can define another index of diversity by normalizing the diversity between the subpopulations by its (best) upper bound. If the concave function used in the definition of J_D is either the Shannon entropy or the entropy of degree α , then the index of diversity can be defined as follows:

$$(5.2) \quad G_D^* = \begin{cases} \frac{J_D(\{P_i\}, \{x_i\}) (2^{1-\alpha} - 1)}{n^{1-\alpha} - 1} & \text{if } D = H_\alpha \\ \frac{J_D(\{P_i\}, \{x_i\})}{\log_2 m} & \text{if } D = H. \end{cases}$$

In summary, we have shown that the diversity between subpopulations (or the generalized Jensen difference) induced by either the Shannon entropy or the entropy of degree α can never exceed the corresponding entropies of the prior probabilities of subpopulations. The second order generalized Jensen difference (which is also interaction between factor 1 and 2) is bounded from above by the sum of the entropies of prior distributions of the factor 1 and 2.

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