## Kybernetika

Jiří Anděl; Václav Dupač<br>Extrapolations in non-linear autoregressive processes

Kybernetika, Vol. 35 (1999), No. 3, [383]--389
Persistent URL: http://dml.cz/dmlcz/135294

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# EXTRAPOLATIONS IN NON-LINEAR AUTOREGRESSIVE PROCESSES 

Jirí Anděl and Václav Dupač

We derive a formula for $m$-step least-squares extrapolation in non-linear $\mathrm{AR}(p)$ processes and compare it with the naïve extrapolation. The least-squares extrapolation depends on the distribution of white noise. Some bounds for it are derived that depend only on the expectation of white noise. An example shows that in general case the difference between both types of extrapolation can be very large. Further, a formula for least-squares extrapolation in multidimensional non-linear $\operatorname{AR}(p)$ process is derived.

## 1. INTRODUCTION

Let $e_{1}, e_{2}, \ldots$ be i.i.d. random variables with finite second moment. Assume that $e_{t}$ has a density $h$. Let $X_{0}, X_{-1}, \ldots, X_{-p+1}$ be random variables independent of $\left\{e_{t}, t \geq 1\right\}$. Define $\gamma=E e_{t}$ and

$$
\begin{equation*}
X_{t}=\lambda\left(X_{t-1}, \ldots, X_{t-p}\right)+e_{t}, \quad t \geq 1 \tag{1.1}
\end{equation*}
$$

where $\lambda$ is a Borel measurable (generally a non-linear) function such that $E X_{t}^{2}<$ $<\infty$. Then $\left\{X_{t}, t \geq 1\right\}$ is called non-linear autoregressive process of order $p$, briefly $\operatorname{NLAR}(p)$.

Assume that for a $t \geq 1$ the variables $\left\{X_{t-s}, s \geq 0\right\}$ are given and $X_{t+m}$ is to be extrapolated. The naïve extrapolation $X_{t+m \mid t}^{*}$ is defined as follows. Let $X_{t+m \mid t}^{*}=$ $X_{t+m}$ for $m \leq 0$ and

$$
X_{t+m \mid t}^{*}=\lambda\left(X_{t+m-1 \mid t}^{*}, \ldots, X_{t+m-p \mid t}^{*}\right)+\gamma, \quad m \geq 1
$$

Denote $\boldsymbol{z}=\left(z_{1}, \ldots, z_{p}\right)^{\prime}$ and introduce functions

$$
H_{m}(\boldsymbol{z})= \begin{cases}z_{-m+1} & \text { for } m=0,-1, \ldots,-p+1  \tag{1.2}\\ \lambda\left[H_{m-1}(\boldsymbol{z}), \ldots, H_{m-p}(\boldsymbol{z})\right]+\gamma & \text { for } m \geq 1\end{cases}
$$

Then

$$
X_{t+m \mid t}^{*}=H_{m}\left(X_{t}, X_{t-1}, \ldots, X_{t-p+1}\right)
$$

Note that $H_{m}(\boldsymbol{z})=H_{1}\left[H_{m-1}(\boldsymbol{z}), \ldots, H_{m-p}(\boldsymbol{z})\right]$.

The least-squares (LS) extrapolation $\hat{X}_{t+m \mid t}$ of the variable $X_{t+m}$ given $\left\{X_{t-s}\right.$, $s \geq 0\}$ is defined by

$$
\hat{X}_{t+m \mid t}=E\left(X_{t+m} \mid X_{t}, X_{t-1}, \ldots, X_{-p+1}\right)
$$

Since the model (1.1) is Markovian of the $p$ th order, we see that

$$
\hat{X}_{t+m \mid t}=E\left(X_{t+m} \mid X_{t}, X_{t-1}, \ldots X_{t-p+1}\right)
$$

Then the variable $\hat{X}_{t+m \mid t}$ can be written in the form

$$
\hat{X}_{t+m \mid t}=K_{m, t}\left(X_{t}, X_{t-1}, \ldots X_{t-p+1}\right)
$$

where $K_{m, t}$ is a function.

## 2. PROPERTIES OF EXTRAPOLATIONS

Calculation of LS extrapolation can be based on the following theorem.

Theorem 2.1. Functions $K_{m, t}, m \geq 0, t \geq 0$, are independent of $t$. They satisfy, with subscript $t$ already dropped, the relation

$$
\begin{equation*}
K_{m}\left(z_{1}, \ldots, z_{p}\right)=\int K_{m-1}\left(w, z_{1}, \ldots, z_{p-1}\right) h\left[w-\lambda\left(z_{1}, \ldots, z_{p}\right)\right] \mathrm{d} w, \quad m \geq 1 \tag{2.1}
\end{equation*}
$$

Proof. Introduce $p$-vectors $\boldsymbol{X}_{\boldsymbol{t}}=\left(X_{t}, \ldots, X_{t-p+1}\right)^{\prime}, \boldsymbol{t} \geq 0, \boldsymbol{u}=(1,0, \ldots, 0)^{\prime}$ and a transformation

$$
T(\boldsymbol{z})=\left(\lambda(\boldsymbol{z}), z_{1}, z_{2}, \ldots, z_{p-1}\right)^{\prime}, \quad \boldsymbol{z} \in \mathbb{R}^{p}
$$

We have

$$
\boldsymbol{X}_{t+1}=T\left(\boldsymbol{X}_{t}\right)+e_{t+1} \boldsymbol{u}, \quad t \geq 0
$$

hence, $\left\{\boldsymbol{X}_{t}\right\}$ is a discrete parameter $p$-dimensional Markov process with the initial distribution of $\boldsymbol{X}_{0}$ and with a stationary transition distribution function

$$
F(\boldsymbol{z}, \boldsymbol{y})=H\left(y_{1}-\lambda(\boldsymbol{z})\right) I\left(y_{2}-z_{1}\right) \ldots I\left(y_{p}-z_{p-1}\right)
$$

where $H$ is the distribution function of $e_{1}$ and

$$
I(y)=\left\{\begin{array}{lll}
1 & \text { for } \quad y \geq 0 \\
0 & \text { for } & y<0
\end{array}\right.
$$

Denote by $F^{(m)}(\boldsymbol{z}, \boldsymbol{y})$ the $m$-step transition distribution function of the process. As $K_{m, t}(\boldsymbol{z})=E\left(\boldsymbol{X}_{t+m}^{\prime} \boldsymbol{u} \mid \boldsymbol{X}_{t}=\boldsymbol{z}\right)$, the independence of $t$ follows. Further,

$$
\begin{gathered}
K_{m}(\boldsymbol{z})=\int_{\mathbb{R}^{p}} \boldsymbol{y}^{\prime} \boldsymbol{u} F^{(m)}(\boldsymbol{z}, \mathrm{d} \boldsymbol{y})=\int_{\mathbb{R}^{p}} \int_{\mathbb{R}^{\boldsymbol{p}}} \boldsymbol{y}^{\prime} \boldsymbol{u} F(\boldsymbol{z}, \mathrm{~d} \boldsymbol{w}) F^{(m-1)}(\boldsymbol{w}, \mathrm{d} \boldsymbol{y}) \\
= \\
\int_{\mathbb{R}^{p}} K_{m-1}(\boldsymbol{w}) F(\boldsymbol{z}, \mathrm{~d} \boldsymbol{w})=\int_{-\infty}^{\infty} K_{m-1}\left(w_{1}, z_{1}, z_{2}, \ldots, z_{p-1}\right) h\left(w_{1}-\lambda(\boldsymbol{z})\right) \mathrm{d} w_{1},
\end{gathered}
$$

proving thus (2).

Theorem 1 for $p=1$ and with an unnecessary stationarity assumption of the process $\left\{X_{t}\right\}$ is given in Tong [3].

It is clear that $K_{0}(\boldsymbol{z})=z_{1}$. Moreover, it follows immediately from our definitions that $H_{1}(\boldsymbol{z})=K_{1}^{\prime}(\boldsymbol{z})=\lambda(\boldsymbol{z})+\gamma$. If $m \geq 2$ then $K_{m}^{\prime} \neq H_{m}$ generally holds, since

$$
\begin{aligned}
H_{2}(\boldsymbol{z}) & =\lambda\left[\lambda(\boldsymbol{z})+\gamma, z_{1}, \ldots, z_{p-1}\right]+\gamma \\
K_{2}(\boldsymbol{z}) & =\int \lambda\left[y+\lambda(\boldsymbol{z}), z_{1}, \ldots, z_{p-1}\right] h(y) \mathrm{d} y+\gamma
\end{aligned}
$$

In special cases, $K_{m}^{\prime}=H_{m}$ holds even for some $m \geq 2$. A typical situation is described in the following theorem.

Theorem 2.2. Let $p \geq 2$ and let

$$
\begin{equation*}
X_{t}=b X_{t-1}+\varphi\left(X_{t-2}, \ldots, X_{t-p}\right)+e_{t} \tag{2.2}
\end{equation*}
$$

where $\varphi$ is a Borel measurable function. Then $K_{2}(\boldsymbol{z})=H_{2}(\boldsymbol{z})$ holds for all $\boldsymbol{z}$.
Proof. In our model (2.2) we have

$$
K_{1}(\boldsymbol{z})=H_{1}(\boldsymbol{z})=b z_{1}+\varphi\left(z_{2}, \ldots, z_{p}\right)+\gamma .
$$

Using (2.1) we get

$$
\begin{aligned}
K_{2}^{\prime}(z) & =\int\left[b y+\varphi\left(z_{1}, \ldots, z_{p-1}\right)+\gamma\right] h\left[y-b z_{1}-\varphi\left(z_{2}, \ldots, z_{p}\right)\right] \mathrm{d} y \\
& =\int\left[b u+b^{2} z_{1}+b \varphi\left(z_{2}, \ldots, z_{p}\right)+\varphi\left(z_{1}, \ldots, z_{p-1}\right)+\gamma\right] h(u) \mathrm{d} u \\
& =b \gamma+b^{2} z_{1}+b \varphi\left(z_{2}, \ldots, z_{p}\right)+\varphi\left(z_{1}, \ldots, z_{p-1}\right)+\gamma
\end{aligned}
$$

Since

$$
X_{t+2}=b X_{t+1}+\varphi\left(X_{t}, \ldots, X_{t-p+2}\right)+e_{t+2}
$$

and the naïve extrapolation of $X_{t+1}$ is

$$
X_{t+1 \mid t}^{*}=b X_{t}+\varphi\left(X_{t-1}, \ldots, X_{t-p+1}\right)+\gamma
$$

we get

$$
\begin{aligned}
X_{t+2 \mid t}^{*} & =b X_{t+1 \mid t}^{*}+\varphi\left(X_{t}, \ldots, X_{t-p+1}\right)+\gamma \\
& =b \gamma+b^{2} X_{t}+b \varphi\left(X_{t-1}, \ldots, X_{t-p+1}\right)+\varphi\left(X_{t}, \ldots, X_{t-p+1}\right)+\gamma
\end{aligned}
$$

From this formula for $X_{t+2 \mid t}^{*}$ it is clear that $K_{2}(z)=H_{2}(z)$.
Consider again the general model (1.1). It is clear that the naïve extrapolation depends on the white noise $\left\{e_{t}\right\}$ only through its expectation $\gamma$ whereas the LS extrapolation depends on the complete distribution of $e_{t}$. It can be of some use to have some bounds for LS extrapolation such that they depend also only on $\gamma$. We derive such results for a class of functions $\lambda(\boldsymbol{z})$. First of all, we introduce auxiliary assertions.

Lemma 2.3. Assume that $f(\boldsymbol{z})$ is a concave function such that it is non-decreasing in each variable. Let $g_{1}(\boldsymbol{z}), \ldots, g_{p}(\boldsymbol{z})$ be concave functions. Then the function $f\left[g_{1}(\boldsymbol{z}), \ldots, g_{p}(\boldsymbol{z})\right]$ is concave.

Proof. The assertion can be proved directly.
Lemma 2.4. Let $\xi$ be a random variable such that $a \leq \xi \leq b$ where $-\infty<a<$ $b<\infty$. Define $p=(E \xi-a) /(b-a), q=1-p$. If $\phi$ is a convex function on $[a, b]$ then

$$
\begin{equation*}
E \phi(\xi) \leq q \phi(a)+p \phi(b) \tag{2.3}
\end{equation*}
$$

Proof. See Kall and Wallace [2], p. 168.
The inequality (2.3) is called Edmundson-Madansky upper bound.

Lemma 2.5. Let $\lambda(z)$ be a concave function. Then the functions $H_{m}(z), m \geq$ $-p+1$, are concave and non-decreasing in each variable.

Proof. Theorem can be proved by complete induction using Lemma 2.3.

Lemma 2.6. Let

$$
\begin{aligned}
U_{1}(\boldsymbol{z}) & =H_{1}(\boldsymbol{z}) \\
U_{m}(\boldsymbol{z}) & =U_{m-1}\left[\gamma+\lambda(\boldsymbol{z}), z_{1}, \ldots, z_{p-1}\right], \quad m \geq 2
\end{aligned}
$$

If $\lambda(z)$ is a concave function non-decreasing in each variable then the functions $U_{m}(z), m \geq 1$, are concave.

Proof. The assertion follows from Lemma 2.3 using complete induction.
By the way, it is easy to see that $U_{2}(\boldsymbol{z})=H_{2}(\boldsymbol{z})$.
Theorem 2.7. Let $\lambda(z)$ be a concave function non-decreasing in each variable. Then $K_{m}(\boldsymbol{z}) \leq U_{m}(\boldsymbol{z}), m \geq 1$.

Proof. We use complete induction. We know already that $K_{1}(\boldsymbol{z})=H_{1}(\boldsymbol{z})=$ $U_{1}(z)$. Let $m=2$. Our assumptions ensure that the function $r(y)=\lambda[y+$ $\left.\lambda(z), z_{1}, \ldots, z_{p-1}\right]$ is concave for arbitrary fixed $\boldsymbol{z}$. Jensen inequality gives

$$
\begin{aligned}
K_{2}(\boldsymbol{z}) & =\gamma+\int \lambda\left[y+\lambda(\boldsymbol{z}), z_{1}, \ldots, z_{p-1}\right] h(y) \mathrm{d} y \\
& \leq \gamma+\lambda\left[\gamma+\lambda(\boldsymbol{z}), z_{1}, \ldots, z_{p-1}\right] \\
& =H_{2}(\boldsymbol{z})=U_{2}(\boldsymbol{z})
\end{aligned}
$$

For $m>2$ it follows from the induction assumption, Lemma 2.6 and Jensen inequality that

$$
\begin{aligned}
K_{m}(\boldsymbol{z}) & =\int K_{m-1}\left[y+\lambda(\boldsymbol{z}), z_{1}, \ldots, z_{p-1}\right] h(y) \mathrm{d} y \\
& \leq \int U_{m-1}\left[y+\lambda(\boldsymbol{z}), z_{1}, \ldots, z_{p-1}\right] h(y) \mathrm{d} y \\
& \leq U_{m-1}\left[\gamma+\lambda(\boldsymbol{z}), z_{1}, \ldots, z_{p-1}\right]=U_{m}(\boldsymbol{z})
\end{aligned}
$$

Theorem 2.8. Let $a \leq e_{t} \leq b$ where $-\infty<a<b<\infty$. Define $p=(\gamma-a) /(b-a)$, $q=1-p$ and

$$
\begin{aligned}
J_{1}(\boldsymbol{z}) & =\lambda(\boldsymbol{z}) \\
J_{m}(\boldsymbol{z}) & =q J_{m-1}\left[a+\lambda(\boldsymbol{z}), z_{1}, \ldots, z_{p-1}\right]+p J_{m-1}\left[a+\lambda(\boldsymbol{z}), z_{1}, \ldots, z_{p-1}\right], \quad m \geq 2 \\
L_{m}(\boldsymbol{z}) & =\gamma+J_{m}(\boldsymbol{z}), \quad m \geq 1
\end{aligned}
$$

If $\lambda(z)$ is a concave function non-decreasing in each variable then $J_{m}, m \geq 1$, are concave functions and $L_{m}(z) \leq K_{m}(z)$ for $m \geq 2$.

Proof. It is clear that $J_{m}(\boldsymbol{z})$ are concave functions. From Lemma 2.4 we have

$$
K_{2}(\boldsymbol{z})=\gamma+\int_{a}^{b} \lambda\left[y+\lambda(\boldsymbol{z}), z_{1}, \ldots, z_{p-1}\right] h(y) \mathrm{d} y \geq \gamma+J_{2}(\boldsymbol{z})=L_{2}(\boldsymbol{z}) .
$$

Similarly, by complete induction we obtain for $m \geq 3$ that

$$
\begin{aligned}
K_{m}(\boldsymbol{z}) & =\int_{a}^{b} K_{m-1}\left[y+\lambda(\boldsymbol{z}), z_{1}, \ldots, z_{p-1}\right] h(y) \mathrm{d} y \\
& \geq \gamma+\int_{a}^{b} J_{m-1}\left[y+\lambda(\boldsymbol{z}), z_{1}, \ldots, z_{p-1}\right] h(y) \mathrm{d} y \\
& \geq \gamma+J_{m}(\boldsymbol{z})=L_{m}(\boldsymbol{z}) .
\end{aligned}
$$

## 3. AN EXAMPLE

Let $v>0$. Assume that $e_{t} \sim R(0, v)$ and define

$$
X_{t}=\sqrt{X_{t-1} X_{t-2}}+e_{t}, \quad t \geq 1
$$

where $X_{0}$ and $X_{1}$ are non-negative random variables. Then $\gamma=v / 2$. From $X_{t+1}=$ $\sqrt{X_{t} X_{t-1}}+e_{t+1}, X_{t+2}=\sqrt{X_{t+1} X_{t}}+e_{t+2}$ we get

$$
H_{1}\left(z_{1}, z_{2}\right)=\sqrt{z_{1} z_{2}}+\gamma, \quad H_{2}\left(z_{1}, z_{2}\right)=\sqrt{\left(\sqrt{z_{1} z_{2}}+\gamma\right) z_{1}}+\gamma
$$

It was already mentioned that $K_{1}\left(z_{1}, z_{2}\right)=H_{1}\left(z_{1}, z_{2}\right)$. From (2.1) we obtain

$$
K_{2}\left(z_{1}, z_{2}\right)=\gamma+\frac{2}{3 v} \sqrt{z_{1}}\left[\left(v+\sqrt{z_{1} z_{2}}\right)^{3 / 2}-\left(z_{1} z_{2}\right)^{3 / 4}\right]
$$

For $v=10$ the functions $H_{1}, H_{2}, K_{2}$, and $D_{2}=K_{2}-H_{2}$ are plotted in Figures 1, 2, 3, and 4 , respectively.


Fig. 1.


Fig. 3.


Fig. 2.


Fig. 4.

Consider the difference

$$
\delta(v)=K_{2}(1,1)-H_{2}(1,1)=\sqrt{1+\frac{v}{2}}-\frac{2}{3 v}\left[(v+1)^{3 / 2}-1\right] .
$$

It is clear that $\delta(v) \rightarrow \infty$ as $v \rightarrow \infty$. This is an elementary example that the difference between the least-squares and the naïve extrapolations can be arbitrary large. A similar example for an NLAR(1) process is introduced in Anděl [1].

## 4. MULTIDIMENSIONAL NLAR $(p)$ PROCESSES

Now, let $\boldsymbol{e}_{1}, \boldsymbol{e}_{2}, \ldots$ be i.i.d. random $q$-vectors with finite second-order moments. Assume that $\boldsymbol{e}_{t}$ has a density $h$. Define $\boldsymbol{\gamma}=E \boldsymbol{e}_{\boldsymbol{t}}$. Let $\boldsymbol{X}_{0}, \boldsymbol{X}_{-1}, \ldots, \boldsymbol{X}_{-p+1}$ be random $q$-vectors, independent of $\left\{e_{t}, t \geq 1\right\}$. Let $\lambda$ be a Borel measurable function from $\mathbb{R}^{p q}$ into $\mathbb{R}^{q}$. For $q$-vectors $\boldsymbol{z}_{1}, \ldots, \boldsymbol{z}_{p}, \lambda\left(\boldsymbol{z}_{1}, \ldots, \boldsymbol{z}_{p}\right)$ or $\lambda(\mathcal{Z})$ will denote the function value of $\lambda$ at point $\mathcal{Z}=\operatorname{vec}\left(z_{1}, \ldots, z_{p}\right)$. Define

$$
\boldsymbol{X}_{t}=\lambda\left(\boldsymbol{X}_{t-1}, \ldots, \boldsymbol{X}_{t-p}\right)+\boldsymbol{e}_{t}, \quad t \geq 1
$$

and assume that the second order moments of $\boldsymbol{X}_{\boldsymbol{t}}$ are finite for all $t \geq 1$. Thus $\left\{\boldsymbol{X}_{t}, t \geq 1\right\}$ is a $q$-dimensional $\operatorname{NLAR}(p)$ process.

Introduce functions $H_{m}: \mathbb{R}^{p q} \rightarrow \mathbb{R}^{q}$ by

$$
H_{m}(\mathcal{Z})= \begin{cases}\boldsymbol{z}_{-m+1} & \text { for } m=0,-1, \ldots,-p+1 \\ \lambda\left[H_{m-1}(\mathcal{Z}), \ldots, H_{m-p}(\mathcal{Z})\right]+\boldsymbol{\gamma} & \text { for } m \geq 1\end{cases}
$$

Then $\boldsymbol{X}_{t+m \mid t}^{*}=H_{m}\left(\boldsymbol{X}_{t}, \ldots, \boldsymbol{X}_{t-p+1}\right)$ is the naïve extrapolation of $\boldsymbol{X}_{t+m}$ given $\boldsymbol{X}_{\boldsymbol{t}}, \boldsymbol{X}_{\boldsymbol{t}-1}, \ldots$

Theorem 4.1. Define functions $K_{m}: \mathbb{R}^{p q} \rightarrow \mathbb{R}^{q}$ by

$$
\begin{aligned}
K_{0}^{\prime}\left(\boldsymbol{z}_{1}, \ldots, \boldsymbol{z}_{p}\right) & =\boldsymbol{z}_{1} \\
K_{m}^{\prime}\left(\boldsymbol{z}_{1}, \ldots, \boldsymbol{z}_{p}\right) & =\int K_{m-1}\left(\boldsymbol{w}, \boldsymbol{z}_{1}, \ldots, \boldsymbol{z}_{p-1}\right) h\left[\boldsymbol{w}-\lambda\left(\boldsymbol{z}_{1}, \ldots, \boldsymbol{z}_{p}\right)\right] \mathrm{d} \boldsymbol{w}, \quad m \geq 1
\end{aligned}
$$

Then

$$
\hat{\boldsymbol{X}}_{t+m \mid t}=K_{m}\left(\boldsymbol{X}_{t}, \ldots, \boldsymbol{X}_{t-p+1}\right)
$$

is the least-squares extrapolation of $\boldsymbol{X}_{t+m}$ given $\boldsymbol{X}_{t}, \boldsymbol{X}_{t-1}, \ldots$.
Proof. Introduce $\mathcal{X}_{\boldsymbol{t}}=\operatorname{vec}\left(\boldsymbol{X}_{\boldsymbol{t}}, \ldots, \boldsymbol{X}_{t-p+1}\right)$ and

$$
u=\left(\begin{array}{c}
I \\
\cdots \\
\mathbf{0}
\end{array}\right)
$$

where $I$ is the $q \times q$ identity matrix and 0 the $(p-1) q \times q$ zero matrix. With $T(\mathcal{Z})=\operatorname{vec}\left[\lambda(\mathcal{Z}), z_{1}, \ldots, \boldsymbol{z}_{p-1}\right], \mathcal{Z} \in \mathbb{R}^{p q}$, we have $\mathcal{X}_{t+1}=T\left(\mathcal{X}_{t}\right)+\boldsymbol{u} \boldsymbol{e}_{t+1}$. The rest of the proof is quite the same as that of Theorem 2.1.

## ACKNOWLEDGEMENT

The research was supported by grant 201/97/1176 from the Grant Agency of the Czech Republic.
(Received February 2, 1998.)

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